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Students Ranking, Based on their Abilities on Objective Type Test:
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ABSTRACT

The focus of this paper is to investigate the idea of differential ability for students ranking on a multiple choice test in the subject of physics at secondary level. The weakness of Classical Test Theory (CTT) in measuring the ability with continuity and consequently its ranking ability of students on the basis of ability has been highlighted in this paper. In CTT, a student attempting a difficult question and an easy question get equal credit which is not the case in Item Response Theory (IRT). Moreover in CTT two students with equal raw score have the same ranking while in IRT they have different ranking, making the job of policy maker easier to take decision. Two contemporary approaches, CTT and IRT were compared on their suitability for ranking and measuring true ability on teacher made test. The sample was 400, 9th grade students taken randomly from a variety of population in Pakistan. A content valid test of 80 multiple choice items was used as instrument. This attempt is an illustrative example that this problem can be overcome by using of Item Response theory in measurement. The implication of this work for the teachers is to give more stress on teaching cognitive skills rather than knowledge and for the policy maker to evaluate the students on the basis of their cognitive skills achievement to award scholarship and recruitment.

Keywords: Ranking of Students, Item Response Theory, Students Ability, Classical Test Theory

INTRODUCTION

Assessment is an indispensable part of educational process. A major purpose of assessment in educational settings is to measure students' achievement in order to make a variety of decisions based on students' performance like to know their present level of learning and to what extent are they ready for next learning experiences? "Advances in technology and the growing presence of computers in assessment provide tremendous opportunity to explore new ways to improve the quality of assessment data" (Klein & Hamilton, 1999). The innovation in testing techniques like computer based administration of tests enables us to collect additional information related to the interaction between each individual examinee and a single item on the test. In theory of measurement in education and psychology there are two contemporary approaches namely Classical Test Theory (CTT) and Item Response Theory (IRT). Both are used to measure a sample of behaviour and a numerical value is assign to the behaviour for quantification. In CTT the number of correct score is often taken as ability. Moreover CTT measurement is holistic in nature and is based on the test as a whole while in IRT item

is the unit of assessment. Many studies have been conducted to investigate the comparability of items and person statistic. Theoretically CTT is simple and easy to apply that is why its test statistics are still commonly used in test construction process, however many researchers have questioned their utility in the modern era. Hambleton and Jones (1993) also expressed their reservation about the use of classical test theory estimators by saying that “classical item statistics such as item difficulty and item discrimination (i.e., point biserial correlations) and test statistics such as test reliability are dependent on the examinee sample in which they are obtained”. Similarly Fan (1998), Cantrell (1997) and Henson (1999) has summarized and noted this problem as the estimators coming from CTT are circular dependant i.e. the items parameter depends upon the examinee and the abilities of examinees are function of items parameter.) This circular dependency in the case of easy test can exaggerate the ability estimates of the students and difficult test can do the reverse job by underestimating the abilities of examinees. Therefore, it is difficult to generalize the classical test theory estimators across populations especially when they are at variance in abilities. Courville (in Traub and Rowley, 1991) wrote that classical test reliability shows the quality of a set of test scores and hence, reliability is dependent on characteristics of the group of students who take the test, in addition to being dependent on characteristics administration of the test. Comparison of performance of different examinees is another noteworthy limitation of CTT because the examinees must be given either the same or parallel items. The problem is further accentuated by a third limitation of classical test theory in that parallel forms are difficult to achieve. A fourth problem of classical test theory as reported by Courville (in Hambleton & Swaminathan, 1985) is that it provides no basis for determining how an examinee might perform when confronted with a test item. Finally, the theoretical assumption that the measurement error is the same for all examinees is another limitation of CTT which may lead to inappropriate ability measures.

Due to these criticisms there is a trend for shifting the focus to the Item Response Theory among test developers and other stake holders. Furthermore the appropriateness of both frameworks for ranking of students on the basis of their ability is an open question which we have addressed in this study.

ITEM RESPONSE THEORY: Item response theory (IRT) is, for some researchers, the answer to the limitations of classical test theory as stated by Courville (2004, p.44). Item response theory (IRT) looks at the examinee’s performance by using item as the unit of assessment. Cantrell (1999), Hambleton & Swaminathan, (1985) and Henard (2000) consider IRT as a modelling technique that tries to describe the relationship between an examinee’s test performance and the latent trait underlying the performance.

There are two general factors in measurement while using CTT approach, an observed response (X) i.e. scored obtained by the students on a particular task and a true ability (T) which is the real potential in a student. This relationship in theoretical model of CTT can be written as $X=T+E$ where E is random error of measurement . IRT in comparison to CTT is based upon a family of the mathematical models. Thus, both models are liable to mislead because they are dependent on the assumptions a researcher is putting forward while working with given data (Hambleton & Swaminathan, 1985).

Hambleton and Swaminathan (1985) have pointed out the following four characteristics of an item response model. (1) An IRT model must give specification the relationship between the measured score and the underlying unobservable construct. (2) the model must provide a way to estimate scores on the ability. (3) The examinee’s scores will be the basis for the estimation of the underlying unobservable construct. (4) this model assumes that the performance of an examinee is completely predictable or can be explained from one or more abilities.

Three models popularly used in IRT are 1-parameter, which is

$$P(\theta) = \frac{e^{a(\theta)}}{1 + e^{a(\theta)}}$$

Where $P(\theta)$ is ability of a student and $a(\theta)$ is difficulty level of item and $e=2.73$ discrimination index $b(\theta)$ is taken as 1 in this model. In 1-parameter model discrimination is taken as 1 and this may not be

of great utility where sharp measurement is required e.g. students with equal raw score will have equal IRT score and thus may fail to produce ranking. 2-parameter Model which is

$$P_i(\theta) = \frac{e^{a_i(\theta - b_i)}}{1 + e^{a_i(\theta - b_i)}}$$

In 2-parameter model both discrimination and difficulty of items are taken into account which enables us to differentiate between the abilities of person with equal raw score, similarly 3-parameter model involves another variable i.e. chance factor c_i for item "i" in attempting an item.

$$P(\theta) = c_i + \frac{e^{a(\theta - b)}}{1 + e^{a(\theta - b)}}$$

When the two- and three-parameter item characteristic curve models are used, an examinee's ability estimate depends upon the particular pattern of item responses rather than the raw score. Under these models, examinees with the same item response pattern will obtain the same ability estimate. Thus, examinees with the same raw score could obtain different ability estimates if they answered different items correctly. (Baker, 2000, p.136)

This is one of the reasons that we opt to use the 2-parameter model instead of 1-parameter model. Estimation of ability focuses on an individual student's responses that give maximum information. Only those items are used to estimate students' ability which have difficulty level of 50% which means at this level the students are 50% likely to get the item right and 50% likely to get the item wrong. A good and informative item is one that has moderate difficulty. The test item which is either too difficult or too easy tells us nothing about the students ranking. This is because too much easy item will have the chance to be attempted by almost all of the students without any brainwork and therefore is not feasible to discriminate them on the basis of ability. Similarly, too much difficult will be attempted by none, except guessing, and therefore not feasible for ranking purpose. Cox & Gorsuch (2000) study this property of an item.

Ability is a continuous variable and IRT gives continuous estimates. CTT gives discrete estimates especially in dichotomously made test and may create discrepancies in students' ranking by total raw scores and IRT student ability estimates. This property has been exploited in this work to make the job easier for policy maker in awarding scholarships or admission. The superiority of IRT upon the CTT for the purpose of ranking comes from the fact that on CTT scale many students may have equal raw score and therefore the task of ranking becomes more difficult for policy maker. In IRT, the score is weighted on basis of parameters of item attempted by examinee and therefore it less likely for two examinees two have equal raw score. This case may happen when two or more examinees attempt exactly same items or items with exactly same parameters, but the probability of such a situation is very low. The examinee attempting more difficult items will get higher rank automatically.

Cox and Gorsuch (2000) discussed the point that Students' ability estimates give better measures when using only the items at a level of difficulty at which students are likely to get the item right with probability of 50% , and get the item wrong with the same probability. This feature may create discrepancies in students' ranking in that some students who get higher total raw scores may get lower IRT student ability estimates, and conversely, students who are in higher IRT student ability rank may get lower total raw scores, this discrepancy occurs automatically due to built in mechanism of credit and penalty in IRT approach ; then which set of scores should be accepted. They argued that depending on the total sample size, the IRT student ability estimates is probably the right choice. Even if the sample size is smaller (below 100), which may lead to more error in estimation, ability estimates generated by IRT are likely to be more precise than its counterpart CTT approach. The ability range in IRT estimates is between $-\infty$ to $+\infty$ theoretically but typically they range from +3.0 for student with high abilities on the test to -3.0 for students with low abilities. The two extremes of infinities are for

over fit cases where the students either gives correct answers for all items or gives no correct answer, such cases are omitted from IRT analysis. The difficulty estimates in IRT for items range from +3 to -3. The item with difficulty level +3 and -3 are labelled as "very difficult," and "very easy" respectively.

PURPOSE OF THE STUDY

The main objective of this work is to compare CTT and IRT for their suitability of students ranking on the basis of their scores for awarding scholarships or admission.

METHODS AND MATERIAL

Data Source

Four hundred students selected randomly from both private and public schools in District Malakand of Pakistan, including both genders provided data for this study. A test with 80 multiple choice items was administered.

Sample

Due to tedious nature of IRT analysis without a specialized software and manual marking 100 students were selected randomly for final analysis.

Tools

Matlab software was used for programming to calculate IRT estimates. For this purpose the researcher developed the programming themselves and no specialized software for IRT was used. The program coding can be seen in appendix B.

Description of the Test

The test was developed by the researcher according to table of specification and was validated with the help of subject teacher. Item analysis revealed that the test is consisted of a variety of items from very difficult item with difficulty level of 20 to very easy item with difficulty level of 83.

Item difficulty of each of items in the test was computed by means of the following formula, in which R is the number of students who answered the item correctly and T is the total number of students in the test:

$$\text{Item difficulty} = \frac{R}{T} \times 100$$

To calculate the difficulty level of each item, 27% high achiever and 27% low achiever were taken. It was assumed that the responses of the students in the middle group follow essentially the same pattern. Item discriminating power of a test item refers to the degree to which it discriminates between students with high and low scores. Discrimination power was computed as the difference between the averages percent score of high and low achiever.

RESULT AND DISCUSSION

The test was scored using both approaches and ranking was made on the basis of both CTT and IRT scores. It was observed from the ranking that there is considerable shift of students ranking when it was made on basis of IRT. For example in CTT the top students was with score 72 and second was who scored 68 but in IRT ranking this ranking changed and 68 was on top of the ranking. When the

answer sheet was analysed it was evident that IRT ranking is better because student with score 68 selected wrong choices for easy items and thus got less penalty while students with 72 in CTT relatively could not answer the difficult items and got more penalty consequently lost his ranking. Similarly two students who got 64 score in CTT were placed 4th in CTT ranking while in IRT ranking they were placed 4th and 6th respectively. Four students obtained 60 scores and were all in the same ranking in CTT clearly showing the inability CTT approach to decide which one was better while they were given different ranking in IRT where item was playing the role to decide. Students No.9 who was placed 19th in CTT ranking was given better position in ranking due to the fact that he gave correct choices for most difficult items. During the analysis it was found that only 9 students out of 100 had the same ranking on both scales, 52 students got better ranking while 39 lost their position. This analysis clearly unveils the utility of IRT approach. Analysis of top 15 students has been given in the table below.

S.NO.	IRT	Raw score	IRT Rank	CTT Rank	Difference
1	4.661	68	1	2	-1
2	4.099	72	2	1	1
3	3.823	66	3	3	0
4	3.063	64	4	4	0
5	2.345	63	5	6	-1
6	2.283	64	6	4	2
7	2.066	63	7	6	1
8	2.051	60	8	10	-2
9	1.989	54	9	19	-10
10	1.816	60	11	10	1
11	1.799	58	12	15	-3
12	1.749	60	13	10	3
13	1.861	61	10	8	2
14	1.712	60	14	10	4
15	1.674	57	15	16	-1

The abilities measures form IRT for 25 students out of 100 were between +1 to +3 showing that the ability level was very high for those students as compared to the test difficulty level. While for remaining students the ability level and test difficulty level was comparable.

The correlation between CTT and IRT was found to be 0.95 which indicates a high correlation.

CONCLUSION

For small sample of 100 taken in this study, CTT-based and IRT-based examinee ability estimates were very comparable and highly correlated (0.95), indicating that an analysis of the ability level of individual examinees will lead to similar results across the different measurement theories. This is in accordance with the findings of Courville (2004), Lawson (1991), Fan (1998), Stage (1998), and MacDonald and Paunonen (2002).

The results in this study, based on the differential measurement of IRT, clearly uncover the weakness of the classical test theory in terms of ranking in the cases where it is seriously required. For example in the competition for admission of professional colleges or awarding scholarships like HEC in Pakistan some time a number of candidate obtain equal score and the decision is then left to other measures like score in previous exam or non academic measures like age.

This study has the implications for teacher as well in the sense that teacher should stress in the learning of skill and higher order thinking instead of knowledge based domain which has less value than the former.

The result of this study support the claim of (A. Hotui, 2006) that objective type test can be used to measure high order skills because it was observed that the item which were falling in domain of higher cognitive skill were difficult and thus those students who failed to give correct response got more penalty in terms of losing score . The scores were from both CTT and IRT with a correlation coefficient of 0.95 which supports Courville (2004, p.113) and courville (in Nunnally's, 1979) assertion that when scores obtained by two approaches are correlated they correlate by degree of 0.90 or higher; thus it is really hair splitting to argue about any difference between the two approaches.

RECOMMENDATION

The result of this study is a guide line for policy makers specially those who are engaged in awarding scholarships or giving admission in professional institutions on the basis of test scores to adapt the IRT approach while measuring the abilities which clearly has the advantage of differentiating among students having equal raw score.

This work was carried with a 2-parameter model which takes into account the difficulty and discrimination of items, to exhibit the ability of IRT approach in measuring the ability along continuous line enabling us to distinguish between students with different abilities clearly. The area of possible research is to replicate this study with 3-parameter model where chance factor also comes into play and to see what difference it makes in the ranking of the students.

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Appendix A

The Matlab code for computation of parameters of an item and person.

```
% In the name of Allah, the beneficent, the merciful
%-----Atiq-ur-Rehman Kashmiri-----
% This file is to compute item parameters, given abilities
% and response
clear
load abilities
load response
clear abil resp
for i=1:size(response,2)
    abil=abilities;
    resp=response(:,i);
    save abil abil
    save resp resp
    a=[.1 .1];
    options=optimset('LargeScale','off');
    out(i,:)=fminunc(@likelihood,a,options);
end
function l=likelihood(param)
% This computes likelihood of an item given its parameters
% i.e. difficulty and discrimination
%
% Since the function we maximize take only one argument
% i.e. the variable which we try to maximize
% the other arguments will be called through global command
%
% before we use this file following variable should be specified as global
% 1. abil, the initial guess of abilities of the students
% 2. resp, the response of respondent on the item

load abil
load resp
diff=param(1);
disc=param(2);
for i=1:size(abil)
    arg=exp(disc*abil(i,1)-disc*diff);
    p=arg/(1+arg);
    if resp(i,1)==0
        lik(i,1)=1000*(1-p);
    end
    if resp(i,1)==1
        lik(i,1)=1000*p;
    end
end
l=-log(prod(lik));

% In the name of Allah, the beneficent, the merciful

%This file will compute the revised abilities of students
```

```

% by maximizing likelihood
% given parameters of items and the response
clear
load response
load parameters
clear param resp
for i=1:size(response,1)
    resp=response(i,:);
    save resp resp
    a=0;
    options=optimset('LargeScale','off');
    result(i,1)=fminunc(@likelihood2,a,options);
end
function l=likelihood2(a)

% In the name of Allah, the beneficent, the merciful

% This computes likelihood of an item
% for fixed parameters and
% at given ability level
% the response and the parameters should be
% given as global variables a priori
%
% =====
%
load parameters
load resp

for i=1:size(resp,1)
    param=parameters(i,:);
    diff=param(1);
    disc=param(2);
    arg=exp(disc*a-diff);
    p=arg/(1+arg);
    if resp(i,1)==0
        lo(i,1)=100*(1-p);
    end
    if resp(i,1)==1
        lo(i,1)=100*p;
    end
end
l=-log(prod(lo));

```