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RESEARCH

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# Association between telomere length in peripheral blood leukocytes and risk of ischemic stroke in a Han Chinese population: a linear and non-linear Mendelian randomization analysis

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## Abstract

**Background:** Many contradictory conclusions pertaining to the telomere length in peripheral leukocyte chromosomes as a potential biomarker for ischemic stroke (IS) risk have been reported by the various observational studies in previous years. This study aims to investigate whether the leukocyte telomere length is associated with an increased IS risk or not, based on the Mendelian randomization (MR) approach.

**Methods:** Based on the NHGRI-EBI GWAS Catalog database, the Chinese online genetic database as well as the previous published studies, twelve single nucleotide polymorphisms (SNPs) with minor allele frequency  $\geq 0.05$  were selected and the leukocyte telomere length was measured in 431 first-ever IS patients and 304 healthy controls (quantitative polymerase chain reaction). To explore linear and non-linear effect of telomere length on the IS risk, we performed the linear MR analysis (the inverse-variance weighted method, the maximum likelihood method, and the mode-based estimation method), and the non-linear MR analysis (semiparametric method with three tests for non-linearity, including the quadratic test, Cochran's Q test, and the fractional polynomial test).

**Results:** Two verified SNPs (rs11125529 and rs412658) were chosen as instrumental variables. In linear MR analysis, the adjusted odds ratios and 95% confidence intervals of IS for genetically predicted telomere lengths, based on the two SNPs, were 1.312 (0.979 to 1.759), 1.326 (0.932 to 1.888) and 1.226 (0.844 to 1.781) for the inverse-variance weighted method, the maximum likelihood method, and the mode-based estimation method, respectively. Three tests for nonlinearity failed to reject the null exactly, indicating that the relationship between telomere length and IS risk is unlikely to be non-linear.

**Conclusion:** This MR study based on individual data does not provide strong evidence for a positive linear or non-linear effect of telomere length on the IS risk.

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**Keywords:** Telomere length, Ischemic stroke, Mendelian randomization analysis, Disease biomarker

## Background

In China, stroke is one of the leading causes of death, and about 78% of all strokes are ischemic stroke (IS) [1]. IS is a complex syndrome triggered by cerebral embolism, with its major etiological subtypes being large artery stroke, cardioembolic stroke as well as the small vessel stroke [2]. Chronological age is one of critical factors in IS pathology but is not considered to be a causal factor [3]. This relationship might reflect the impact of biological aging, particularly, in regards to the accumulation of endothelial and vascular damage over time [4]. Telomeres are specialized DNA–protein complexes at linear chromosome ends in eukaryotes, and have been considered as a marker of biological aging at both the cellular and organism level [5, 6]. They safeguard genome stability by preventing DNA degradation and fusion at chromosomal termini [7]. The DNA component that include TTAGGG nucleotide repeats progressively shortens with each cell division. Telomere shortening to a critical length eventually induces cell senescence and programmed cell death, thereby they can be likened to a ‘mitotic clock’ reflecting the amount of cellular turnover within an individual [7]. Furthermore, the telomere attrition rate increased due to oxidative stress and inflammation, both of which are implicated as major mechanisms underlying biological aging [8–10]. Because of its presumed effect on biological aging, telomere length (TL) has been proposed as a potential biomarker for IS [11]. In addition to biological aging, TL shortening has been linked to various risk factors of stroke, including obesity, smoking, alcohol intake, psychological stress and depression [11].

Conventional observational studies have shown inconclusive associations of telomere shortening and the risk of stroke. Consistently, several original studies reported positive association between telomere shortening and the risk of stroke [12–14]. In contrast, evidence also suggests that shorter telomeres were not significantly associated with stroke risk [15–17]. However, meta-analysis have shown that there were less certain results to support the relationship between TL and stroke [18, 19].

Randomized controlled trials of interventions specific to TL have not yet been done in relation to stroke outcomes. In the absence of such trials, Mendelian randomization (MR) analysis focused on genetic instrument can be used to investigate the causality from routinely observational studies [20]. Based on two-sample MR design, three MR analyses have been performed to investigate the association between shorter telomeres and stroke risk and showed contradictory conclusions. First, a MR

study showed that shorter telomeres lead to a marginally significantly decreased odds of stroke in individual-level data [21]. Another two MR studies suggested that TL may be a marker of IS and its subtypes rather than a cause [22, 23]. Two-sample MR typically assume that the exposure-outcome association is linear or log-linear [24]. Therefore, a non-linear TL-IS relationship, such as U-shaped, cannot be detected through this design. To tackle this nonlinearity problem, one-sample MR analysis can be used to explore the non-linear effect relating TL to IS [25]. In this present study, we performed one-sample MR analysis with individual-level data in a Han Chinese population to decipher the linear and non-linear causal role of TL in the IS risk, and to provide insight into the potential mechanisms.

## Methods

### Study population

All participants were recruited from Jidong Oil-field Hospital, Chinese National Petroleum, and Beijing Tiantan Hospital, Capital Medical University, during 2010–2013. A total of 755 participants aged 18 or above were found to be eligible. All participants who met any of the following criteria were excluded from the study: (1) history of mental illness or infectious disease; (2) history of aneurysm caused by cerebral infarction, cerebral haemorrhage or other cerebrovascular diseases, congenital heart disease, acute myocardial infarction, liver disease, renal failure, malignant tumour, chronic obstructive pulmonary disease, rheumatoid arthritis, or other diseases; and (3) pregnant or lactating women. All first-ever IS patients were diagnosed according to the World Health Organization criteria [26]. In this study, 20 participants of non-Han Chinese descent were subsequently excluded. A total of 431 patients with IS and 304 healthy subjects were included in the final analysis.

This study was approved by the Ethics Committee of Capital Medical University, China (No. 2016SY23). This study was in accordance with the principles of the Declaration of Helsinki. All participants provided their written informed consent before taking part in this study.

### Leukocyte telomere length measurement

Blood samples were collected and processed according to the standardized protocol. Following 10 h. of overnight fasting, blood samples were collected by venipuncture in two different tubes containing an anticoagulant and a coagulant respectively. Samples were processed within 8 h. and stored at -80C until further

measurement. Given that whole blood TL is a proxy for tissue-specific TL for many tissues, and blood TL is a proxy for TL in some tissues in epidemiological studies [27], in this study we measured TL in whole blood, but not different types of leukocytes.

Serum samples were used to measure the biochemical indices, and blood cells were used to extract the genomic DNA. Based on automated nucleic acid purification platform (BioTeke, AU1001-32, Beijing, China), genomic DNA was extracted from the 200  $\mu$ l of frozen blood cells using the magnetic bead-based method for concentrating DNA (The Genomic DNA Magnetic Beads Kit, AU18016, BioTeke, Beijing, China). DNA concentration and purity were determined using a Nanodrop (ND-8000, Thermo Scientific, USA) and then normalized to 5 ng/ $\mu$ l in TE buffer (1 mM EDTA, 10 mM Tris-HCl [pH 8.0]). Absolute TL of each chromosome end of the peripheral leukocytes was measured by using a validated quantitative polymerase chain reaction (qPCR) on the method reported by O'Callaghan with modification [28, 29]. A single copy reference gene (*36B4*) which encodes the acidic ribosomal phosphoprotein P0, was used to normalize the quantity of the input DNA. Comparing telomere-to-*36B4* ratio to reference DNA of known TL reflects the absolute length of the telomeres. The master mix was prepared using Applied Biosystems reagents. Primer concentrations were: TelF 200 nM, TelR 200 nM, 36B4F 300 nM, and 36B4R 500 nM. The primer sequences (5'-3') were as follows:

TelF CGGTTTGGTTGGGTTGGGTTGGGTTGGGTTGGGTTGGTTGGGTT

TelR GGCTTGCCCTACCCTTACCCTTACCCTTACCCTTACCCTTACCCTTACCCTT

36B4F CAGCAAGTGGGAAGGTGTAATCC

36B4R CCCATTCTATCATCAACGGGTACAA

Analysis of both telomere and *36B4* in experimental samples and reference DNA were run in triplicates using 5 ng DNA and both telomere and *36B4* amplifications were performed on the same run. A no template control of nuclease-free water was included in each run. In brief, qPCR was performed in 0.1 ml tubes on the Applied Biosystems QuantStudio 3 Real-Time PCR System (Thermo Scientific, USA), with the thermal cycling profile for both telomere and *36B4* amplifications: 2 min at 50 °C, 2 min at 95 °C, followed by 35 cycles of 95 °C for 15 s, 61 °C for 1 min, followed by a melt curve. The no template controls in all runs were no amplification. The melt curve should produce only one single distinct peak. Standard deviation for the Ct-value in replicates were less than 0.5. The inter-coefficients of variance for the Ct-value was less than 5%.

### Genotyping

Single nucleotide polymorphisms (SNPs) were selected from the NHGRI-EBI GWAS Catalog (<https://www.ebi.ac.uk/gwas>) and the online genetic database (<https://www.ensembl.org>) according to reported associations with TL and minor allele frequency (MAF)  $\geq$  0.05. Only one SNP (rs17653722) is reported in population of Chinese descent and other TL-related SNPs were based on the European ancestry. The exclusion of association could be confirmed in online software PhenoScanner (<https://www.phenoscaner.medschl.cam.ac.uk/phenoscaner>). Finally, twelve SNPs (rs11125529, rs6772228, rs10936599, rs7726159, rs9420907, rs17653722, rs3027234, rs8105767, rs409627, rs412658, rs6028466, rs755017) were selected for analysis. SNP genotyping was performed on these SNPs using the Sequenom Massarray iPLEX platform (Agena Bioscience, Inc., San Diego, CA, USA).

### Variables

Demographic characteristics of participants including age, gender, smoking, drinking, height, weight, body mass index (BMI), systolic blood pressure (SBP), diastolic blood pressure (DBP), were collected by questionnaire survey and physical examination. Biochemical data including fasting plasma glucose (FPG), triglycerides (TG), total cholesterol (TC), high-density lipoprotein cholesterol (HDL-C), low-density lipoprotein cholesterol (LDL-C), apolipoprotein A1 (ApoA1), apolipoprotein B (ApoB) were all tested via standard methods in the clinical laboratory of Beijing Tiantan Hospital. TL measurement and SNP genotyping were performed by laboratory personnel according to the standardized protocol.

### Statistical analysis

Demographic and clinical characteristics were represented as the mean  $\pm$  standard deviation (SD) for continuous variables underlying the normal distribution; otherwise, the median (interquartile range) was used. Categorical variables were represented as frequency (percentage). The between-group differences of continuous variables were tested by the *t* test or the Mann-Whitney *U* test. The chi-squared test was used to compare the proportions for categorical variables, and to test for Hardy-Weinberg equilibrium.

The genetic variant or genetic risk score (GRS) acts as an instrumental variable in MR analysis if: (1) they are truly associated with TL; (2) they are independent of other factors (confounders); and (3) they can only influence IS risk via the causal effect of the TL. All association of instrument variables with TL and other risk factors, which were susceptible to reverse association, were

limited to health controls. Principal analyses assumed over-dominant effects (heterosis), with subsidiary analyses of other genetic models (dominant, recessive, co-dominant and additive model).

The  $\beta$  coefficients were obtained from the linear regression model with natural log-transformed TL (ln TL). The percentage difference in TL with risk genotype was obtained from  $[\exp(\beta) - 1] \times 100$ . Then, the linear MR estimates for ln TL on the IS risk were calculated by the inverse-variance weighted method, the maximum likelihood method, and the mode-based estimation method, adjusting for age, sex, and other confounders (smoking status, drinking status, and levels of BMI, SBP, DBP, FPG, TG, TC, HDL-C, LDL-C, ApoA1, ApoB). All results were presented as the odds ratio (OR) of IS per 10% decrease in TL. Three tests for non-linearity of the semiparametric method (the fractional polynomial method and the piecewise linear method) were applied: the quadratic test, which assesses for a linear trend among the localised average causal effect (LACE) estimates, Cochran's Q test, which assesses whether LACE estimates differ more than

would be expected by chance, and the fractional polynomial test, which assesses whether the fractional polynomial model of degree 1 fits LACE estimates better than the linear model [25].

For all analyses, a two-tailed  $P$  value  $< 0.05$  was considered to be statistically significant. All statistical analyses were performed using R version 3.5.3 (R Foundation for Statistical Computing, Vienna, Austria).

## Results

### Participant characteristics

Demographic and clinical characteristics were described in Table 1. Among 735 participants (374 men and 361 women), the median age of the study population was 54 years ( $P_{25}$  44 years,  $P_{75}$  61 years) in all subjects, 47 years (35 to 59 years) in controls, and 57 years (49 to 63 years) in IS patients. Most of the IS patients were older males with smoking and drinking habits and higher weight, BMI, SBP, DBP, TG levels and lower TC, HDL-C, ApoA1, ApoB levels. The levels of TL, height, FPG, LDL-C, and ApoB/ApoA1 ratio were not statistically

**Table 1** Characteristics of the study population

Parameters	Controls ( $n = 304$ )	Cases ( $n = 431$ )	$P$
Age, years	47 (35, 59)	57 (49, 63)	$1.736 \times 10^{-19*}$
Gender, $n$ (%)			$2.404 \times 10^{-4*}$
Male	130 (42.80)	244 (56.60)	
Female	174 (57.20)	187 (43.40)	
Smoke, cigarettes/day, $n$ (%)			$1.227 \times 10^{-17*}$
No	253 (83.20)	248 (57.50)	
1–10	29 (9.50)	41 (9.50)	
> 11	22 (7.20)	142 (32.90)	
Drink, $n$ (%)			$8.893 \times 10^{-5*}$
No	221 (72.70)	252 (58.50)	
Yes	83 (27.30)	179 (41.50)	
TL, kb	6.74 (5.18, 8.43)	6.44 (4.88, 8.93)	0.144
Height, cm	165.0 (159.0, 172.0)	167.0 (160.0, 172.0)	0.324
Weight, kg	66.95 (57.10, 77.58)	70.00 (63.00, 77.80)	0.003*
BMI, kg/m <sup>2</sup>	24.31 (22.08, 27.16)	25.27 (23.07, 27.39)	0.003*
SBP, mmHg	124 (113, 136)	138 (125, 150)	$1.087 \times 10^{-16*}$
DBP, mmHg	79 (72, 88)	82 (76, 91)	$9.324 \times 10^{-7*}$
FPG, mmol/L	5.00 (4.70, 5.50)	5.17 (4.61, 6.00)	0.053
TG, mmol/L	1.26 (0.86, 1.80)	1.34 (1.00, 1.87)	0.017*
TC, mmol/L	4.50 (3.99, 5.05)	4.22 (3.47, 4.99)	$2.451 \times 10^{-4*}$
HDL-C, mmol/L	1.19 (1.01, 1.38)	1.08 (0.91, 1.24)	$5.497 \times 10^{-9*}$
LDL-C, mmol/L	2.53 (2.12, 2.92)	2.46 (1.92, 3.00)	0.232
ApoA1, g/L	1.26 (1.14, 1.36)	1.19 (1.07, 1.29)	$1.727 \times 10^{-9*}$
ApoB, g/L	0.98 (0.85, 1.19)	0.93 (0.73, 1.11)	$1.845 \times 10^{-4*}$
ApoB/ApoA1 ratio	0.79 (0.66, 1.01)	0.77 (0.61, 0.96)	0.198

TL telomere length, BMI body mass index, SBP systolic blood pressure, DBP diastolic blood pressure, FPG fasting plasma glucose, TG triglycerides, TC total cholesterol, HDL-C high-density lipoprotein cholesterol, LDL-C low-density lipoprotein cholesterol, ApoA1 apolipoprotein A1, ApoB apolipoprotein B

\*  $P < 0.05$  is statistically significant

different between groups. There were no statistically significant differences in the genotype distribution of all SNP between the IS patients and the control group (Table 2). Four SNPs with  $MAF < 0.05$  were excluded ( $MAF$ : 0.000 for rs6772228, 0.006 for rs9420907, 0.042 for rs3027234, and 0.002 for rs6028466). SNPs that deviated from Hardy–Weinberg equilibrium were excluded. Other eight SNPs, amongst the healthy controls, were all in accordance with the Hardy–Weinberg equilibrium.

#### Association estimates for individual SNPs

Association of each SNP with  $\ln TL$  in all the assumed genetic models are shown in Additional file 1: Table S1

and Additional file 2: Table S2. To meet MR assumptions basically, two of the SNPs (rs11125529 and rs412658) were used as instrumental variables, and the unweighted GRS were constructed for non-linear MR analysis. By over-dominant model analysis, two SNPs in the control group were found to be associated with decreased TL, with risk genotype difference in  $\ln TL$  of  $-0.108$  (95% confidence interval (95% CI)  $-0.204$  to  $-0.013$ ) for rs11125529,  $-0.089$  ( $-0.176$  to  $-0.001$ ) for rs412658 (Table 3). These two TL-related SNPs were not associated with the conventional risk factors or other biochemical indicators in the control group (Fig. 1). Furthermore, these two SNPs displayed no direct evidence for an

**Table 2 Genotype distributions of SNPs in IS patients and controls**

SNP identifier	Chr	Gene	Genotype	Controls	Cases	<i>P</i>	<i>P</i> <sub>HWE</sub>
rs11125529	2	ACYP2	C/C	202 (66.40%)	305 (70.80%)	0.433	1.000
			C/A	92 (30.30%)	112 (26.00%)		
			A/A	10 (3.30%)	14 (3.20%)		
rs10936599	3	MYNN	T/T	103 (33.90%)	121 (28.10%)	0.240	1.000
			T/C	148 (48.70%)	225 (52.20%)		
			C/C	53 (17.40%)	85 (19.70%)		
rs7726159	5	TERT	C/C	114 (37.50%)	155 (36.00%)	0.252	0.718
			C/A	147 (48.40%)	195 (45.20%)		
			A/A	43 (14.10%)	81 (18.80%)		
rs17653722	12	KRT80	G/G	220 (72.40%)	302 (70.10%)	0.800	0.654
			G/T	76 (25.00%)	117 (27.10%)		
			T/T	8 (2.60%)	12 (2.80%)		
rs8105767	19	LOC112268248	A/A	159 (52.30%)	215 (49.90%)	0.473	0.563
			A/G	125 (41.10%)	177 (41.10%)		
			G/G	20 (6.60%)	39 (9.00%)		
rs409627	19	ZNF676	G/G	129 (42.40%)	189 (43.90%)	0.639	0.304
			G/C	145 (47.70%)	192 (44.50%)		
			C/C	30 (9.90%)	50 (11.60%)		
rs412658	19	ZNF676	C/C	130 (42.80%)	188 (43.60%)	0.596	0.441
			C/T	143 (47.00%)	190 (44.10%)		
			T/T	31 (10.20%)	53 (12.30%)		
rs755017	20	ZBTB46	A/A	102 (33.60%)	163 (37.80%)	0.461	0.907
			A/G	147 (48.40%)	199 (46.20%)		
			G/G	55 (18.10%)	69 (16.00%)		

IS ischemic stroke, SNP single nucleotide polymorphism, Chr chromosome

$P < 0.05$  is statistically significant

$P_{HWE} < 0.05$  indicating that the SNP in control group was not satisfied Hardy–Weinberg equilibrium



**Table 3 Candidate IVs and their association with TL and IS risk under over-dominant model**

	Association with ln TL in controls		Association with IS in all subjects	
	MD (95% CI)	P	OR (95% CI)	P
rs11125529		0.026*		0.203
C/C-A/A	0.000 (0.000, 0.000)		0.000 (0.000, 0.000)	
C/A	-0.108 (-0.204, -0.013)		0.810 (0.580, 1.120)	
rs412658		0.049*		0.428
C/C-T/T	0.000 (0.000, 0.000)		0.000 (0.000, 0.000)	
C/T	-0.089 (-0.176, -0.001)		0.890 (0.660, 1.190)	

TL telomere length, IS ischemic stroke, MD mean difference, 95% CI 95% confidence interval, OR Odds ratio

\* P < 0.05 is statistically significant

individual association with IS risk ( $P > 0.05$ ). The ORs for IS of risk genotype (“short TL”) were 0.810 (0.580, 1.120) for rs11125529, and 0.890 (0.660 to 1.190) for rs412658 (Table 3).

**Association of TL with the IS risk**

In this case–control study, the OR (95% CI) for IS was 0.681 (0.469 to 0.982) adjusted for age and sex (Fig. 2). The association between TL and the IS risk was not statistically significant (OR (95% CI): 0.671 (0.437 to 1.031)) even after further adjusting for the smoking status,

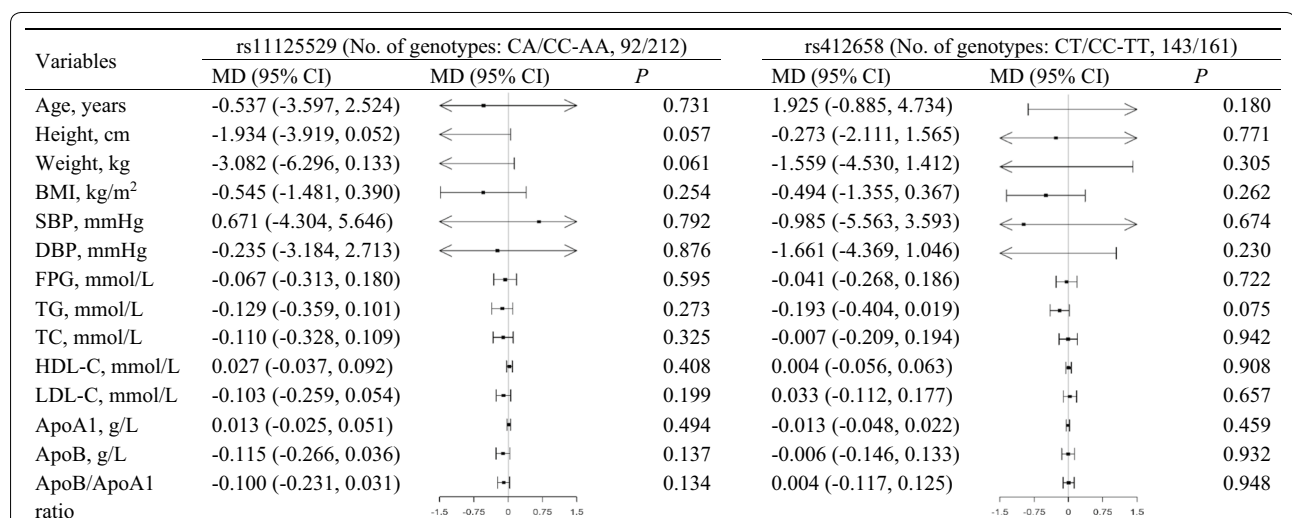
drinking status, and levels of BMI, SBP, DBP, FPG, TG, TC, HDL-C, LDL-C, ApoA1, ApoB.

Using rs11125529 and rs412658 as a proxy, the linear MR analysis provided no evidence of an overall association between genetically predicted TL and IS risk (OR (95% CI): 1.312 (0.979 to 1.759) for the inverse-variance weighted method, 1.326 (0.932 to 1.888) for the maximum likelihood method, and 1.226 (0.844 to 1.781) for the mode-based estimation method), after adjusting for the above-mentioned factors (Fig. 2). Using the unweighted GRS (rs11125529 and rs412658) as instrumental variables, three tests of nonlinearity MR analysis failed to reject the null hypothesis ( $P = 0.069$  for the quadratic test,  $P = 0.126$  for the fractional polynomial test, and  $P = 0.260$  for the Cochran’s Q test), indicating that the effect of telomere attrition on IS risk may not be non-linear. In Fig. 3, the box plot of TL between IS patients and controls across quintiles of the TL was also shown that there was no obvious non-linear TL-IS relationship. Two of five groups by the quintiles of TL, there were statistically significant differences in TL between the IS patients and the controls.

**Discussion**

Based on linear and non-linear MR analysis, we used TL-related SNPs (or unweighted GRS) as proxies to clarify whether TL is causally relevant to IS or not. Our results showed no causal association between the genetically shortened TL and the increased IS risk.

Many epidemiologic studies have examined the relationship between TL and stroke risk, but results were

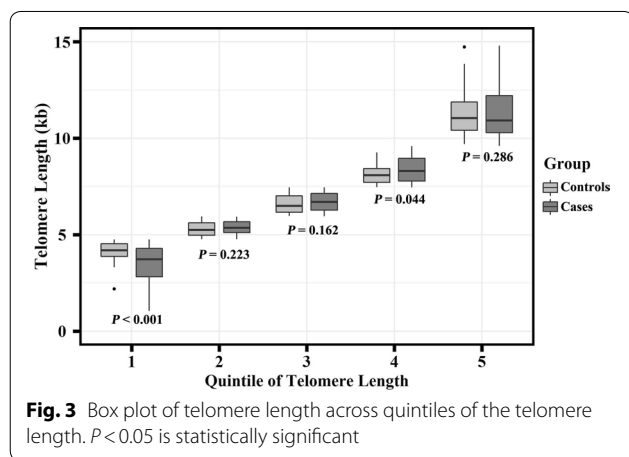


**Fig. 1** Associations of two single nucleotide polymorphisms related to telomere length with various characteristics in individuals free from known ischemic stroke at time of measurement. MD mean difference, 95% CI 95% confidence interval, BMI body mass index, SBP systolic blood pressure, DBP diastolic blood pressure, FPG fasting plasma glucose, TG triglycerides, TC total cholesterol, HDL-C high-density lipoprotein cholesterol, LDL-C low-density lipoprotein cholesterol, ApoA1 apolipoprotein A1, ApoB apolipoprotein B.



	<i>P</i>	OR (95% CI)	
<b>In TL (kb)</b>			
Adjusted for age, sex	<b>0.041*</b>	0.681 (0.469, 0.982)	
Further adjusted†	0.069	0.671 (0.437, 1.031)	
<b>Linear MR analysis†</b>			
Inverse-variance weighted method	0.069	1.312 (0.979, 1.759)	
Maximum likelihood method	0.117	1.326 (0.932, 1.888)	
Mode-based estimation method	0.285	1.226 (0.844, 1.781)	
<b>Non-linear MR analysis†</b>			
The quadratic test	0.069		
The fractional polynomial test	0.126		
The Cochran's <i>Q</i> test	0.260		

**Fig. 2** Estimates of association of telomere length (TL) and genetic prediction telomere with ischemic stroke risk. \**P* < 0.05 is statistically significant. † Corrected for regressions of TL and potential confounding factors, including age, sex, smoking status, drinking status, and levels of body mass index, systolic blood pressure, diastolic blood pressure, fasting plasma glucose, triglycerides, total cholesterol, high-density lipoprotein cholesterol, low-density lipoprotein cholesterol, apolipoprotein A1, apolipoprotein B. OR odds ratio, 95% CI 95% confidence interval, *In TL* Natural log-transformed telomere length, MR mendelian randomization.



**Fig. 3** Box plot of telomere length across quintiles of the telomere length. *P* < 0.05 is statistically significant

not consistent enough. Shorter TL were not significantly associated with IS in the prospective and retrospective studies carried out in the Caucasian [15, 30, 31]. A nested case-control study including 504 case-control pairs of American female nurses showed a non-significant association between relative TL and IS (lowest vs. highest quartile: OR = 0.82, 95% CI 0.52–1.32, *P* = 0.42) [15]. Consistent with findings from the retrospective study, a prospective study (1,136 American adults aged 65 years and older at baseline, average 6.1 years of follow-up, 33 new cases of IS) reported that leukocyte TL change did not associate with IS (1st quartile vs. 4th quartile: relative risk (RR) = 1.61 (95% CI 0.46–5.68; *P* = 0.46) [30]. Similarly, in a cohort study including 4,576 Danish at baseline (10 years of

follow-up, 295 new case of IS), the adjusted RR of IS for TL (4th quartile vs. 1st quartile) were 0.95 (95% CI 0.64–1.40; *P* = 0.60) [31]. In a case-control study carried out in Wuhan (capital of Hubei Province, South China), including 1,309 stroke patients and 1,309 age- and sex-matched control subjects, OR (95% CI) for IS risk were 2.12 (1.62–2.77) compared the first quartile (shortest) to the fourth (longest) quartile [14]. In another case-control study carried out in Shenzhen (Guangdong Province, South China) including 150 cases of IS and 150 siblings of patients free of stroke, shortened TL was independently associated with IS (OR = 4.00, 95% CI 1.28–12.77) [32]. The present MR did not provide strong evidence of causal association between short TL and IS, consistent with findings from the prospective studies and the retrospective studies carried out in the Caucasian, while different to case-control studies performed in the Chinese.

Despite previous retrospective observational studies reporting an TL-stroke association, such a relationship has not been firmly supported by evidence from prospective cohort studies. Subgroup analysis of four meta-analysis (stratified by study design) have shown consistent results that shorter TL were not significantly associated with stroke risk in prospective studies or in studies with a high quality score [18, 19, 33, 34]. In a recent meta-analysis, shorter TL was associated with a significant 11% higher risk for stroke (RR = 1.12, 95% CI 1.05–1.19), although with significant heterogeneity between studies (*I*<sup>2</sup> = 81.1%, *P*<sub>het</sub> < 0.001) [33]. After stratified by study design, the TL-IS relationship was not significance in those prospective studies (RR = 1.03, 95% CI 0.96–1.09)

but remained significant in retrospective studies (RR=1.81, 95% CI 1.54–2.13) [33]. MR analysis which simulates natural experiments based on genetic variants, is considered as an interface between cohort studies and the intervention trials at the evidence level [35]. The inconsistency might be explained by the confounding, reverse causation or recall bias, which might be avoided in MR analysis but not in observational studies.

Three linear MR studies based on two-sample MR design also presented inconsistent results in European ancestry. One MR analysis showed conflicting results which shorter TL was marginally statistically significantly associated with the decreased risk of stroke [21]. Another two linear MR analysis provided no evidence of the linear association between genetically predicted short TL and IS as well as its subtypes [22, 23]. Similarly, we also found null linear relationship between genetically predicted TL and IS risk in a Han Chinese population.

There are several possible explanations for the discrepancy between retrospective and prospective cohort studies. One hypothesis is that TL had causal effect on IS risk under certain circumstances. Difference in factors, such as participant age range, TL measurement technique, and so on, might influence the TL-IS relationship. A previous study indicate that the positive association between short TL and the risk of stroke or post-stroke death might only exist in the seniors population (ranging from 65 to 73 years old), therefore the effect of age might need to be taken into consideration in future studies [11]. However, people aged 65–73 years comprised a relatively low proportion of all participants (9.52%) in our research. Furthermore, although absolute TL measured by qPCR showed a strong correlation ( $r^2=0.75$ ,  $P<0.0001$ ) with the results obtained with terminal restriction fragment analysis (the gold standard for TL measurement). However, slightly difference in TL measurement could affect the TL-IS relationship [28].

Another hypothesis, based on epidemiological evidence, explaining this contradiction is that shorter TL is inversely associated with the risk of IS, which means that TL may be a downstream biological consequence of the IS onset. Figure 3 illustrates that no clear linear or non-linear relationship exists between TL and the risk of IS. However, not all of the differences between the TL and the IS risk were statistically significant within the different TL groups, indicating that there may be a feedback mechanism within certain TL range. Additionally, age is one of the major risk factors of TL and IS, so the effect of age needs to be excluded to prove this hypothesis [7, 11]. Estimates from previous MR studies and our study have avoided the possibility of inverse association to some extent, but more work is still needed to be determined the possible role of TL. For example, bidirectional MR

analysis may be further used to orient the causal direction of TL-IS relationship [36]. Otherwise, in terms of the conflicting results from different research designs, other possibility is that there is no association between TL and IS risk. At present, however, the mechanism research on the relationship of TL and IS is still lacking and lagging, and we cannot rule out the possibility that the existence of certain compensation mechanisms may have affected our results.

For the chief strengths of our study, we explored the possible shape of the potential causal relation between TL and IS risk in a one-sample MR framework using linear and non-linear MR methodology. As a result of the MR analysis, potential reverse causality was eliminated and confounding bias was reduced because genetic instruments were not associated with individual risk factors that may affect results from conventional observational studies. Secondly, although the potential pleiotropic effects were unavoidable in this study, we searched comprehensively from genotype to phenotype to identify the potential pleiotropic effects and further provided possible evidence of the SNP instrument validity that the SNPs have no effect on available confounding factors, to reduce the likelihood of bias due to violation of the instrumental variable analysis. Furthermore, to our knowledge, this is the first linear and non-linear MR study assessing TL in relation to the IS risk in a Han Chinese population. Otherwise, healthy controls were randomly recruited from the general population covering same geographical area, which could decrease the selection bias of the results.

This study also has some limitations. Firstly, MR analysis has stringent assumptions [20]. Completely ruling out potentially pleiotropic effects or an additional biological causal pathway is a challenge for all MR analyses. We are limited by current knowledge and other unavailable confounders, so we cannot exclude the possibility that our estimates are biased by currently unknown pleiotropic effects. Secondly, insufficient statistical power was a common limitation of one-sample MR analysis, and therefore we cannot exclude type II error as an explanation for the null results [37]. Our study does not provide strong evidence for a positive linear or non-linear effect of TL on IS, but does not rule out that genetically predicted TL by unidentified genetic instruments might play a role. Finally, our study was conducted in middle to early late aged participants of Han Chinese descent based on Northern China. Further MR research needs to be explored in a larger and more representative samples, including those from a non-Asian ethnicity.

## Conclusion

In conclusion, although TL is associated with the risk of IS based on the conventional case–control analysis, this one-sample MR study suggests that negative association of TL with the IS risk is unlikely to be linear or non-linear causal. There is a need to identify the specific genetic, biochemical, and environmental biological mechanisms responsible for this association. Further research with larger sample sizes, which will be able to perform stratified analysis by age or other strong risk factors, is necessary to understand the causal pathways underpinning this association.

## Supplementary information

**Supplementary information** accompanies this paper at <https://doi.org/10.1186/s12967-020-02551-1>.

**Additional file 1: Table S1.** Association between SNP genotypes and telomere length under co-dominant and additive model.

**Additional file 1: Table S2.** Association between SNP genotypes and telomere length under dominant, recessive and over-dominant model.

## Abbreviations

ApoA1: Apolipoprotein A1; ApoB: Apolipoprotein B; BMI: Body mass index; DBP: Diastolic blood pressure; FPG: Fasting plasma glucose; GRS: Genetic risk score; HDL-C: High-density lipoprotein cholesterol; IS: Ischemic stroke; LACE: Localised average causal effect; LDL-C: Low-density lipoprotein cholesterol; In TL: Natural log-transformed telomere length; MAF: Minor allele frequency; MR: Mendelian randomization; OR: Odds ratio; qPCR: Quantitative polymerase chain reaction; RR: Relative risk; SBP: Systolic blood pressure; SD: Standard deviation; SNPs: Single nucleotide polymorphisms; TC: Total cholesterol; TG: Triglycerides; TL: Telomere length; 95% CI: 95% confidence interval.

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## Authors' contributions

MS (Manshu Song), YW, and WW contributed to the conception or design of the work. AW, MC, JZ (Jie Zhang) and JZ (Jinxia Zhang) contributed to the acquisition of data. WC, DZ, DL and LW contributed to the analysis and interpretation of data. MS (Manjot Singh), WC, DZ, and IEM contributed to the preparation of the manuscript. All authors read and approved the final manuscript.

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## Availability of data and materials

The datasets used and/or analyzed during the current study are available from the corresponding author on reasonable request.

## Ethics approval and consent to participate

This study was approved by the ethics committees of Capital Medical University, Beijing, China (No. 2016SY23), and the procedures have been performed in accordance with the Declaration of Helsinki. Written informed consent to participate in the study was obtained from all patients enrolled in the study.

## Consent for publication

Not applicable.

## Competing interests

The authors declare that they have no competing interests.

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## References

- Zhao D, Liu J, Wang M, Zhang X, Zhou M. Epidemiology of cardiovascular disease in China: current features and implications. *Nat Rev Cardiol*. 2018;16:203–12.
- Adams HP, Bendixen BH, Kappelle LJ, Biller J, Love BB, Gordon DL, Marsh EE. Classification of subtype of acute ischemic stroke. Definitions for use in a multicenter clinical trial. TOAST. Trial of Org 10172 in Acute Stroke Treatment. *Stroke*. 1993;24:35–41.
- Roy-O'Reilly M, McCullough LD. Age and sex are critical factors in ischemic stroke pathology. *Endocrinology*. 2018;159:3120–31.
- Camicci GG, Savarese G, Akhmedov A, Luscher TF. Molecular mechanism of endothelial and vascular aging: implications for cardiovascular disease. *Eur Heart J*. 2015;36:3392–403.
- Sanders JL, Newman AB. Telomere length in epidemiology: a bio-marker of aging, age-related disease, both, or neither? *Epidemiol Rev*. 2013;35:112–31.
- Fasching CL. Telomere length measurement as a clinical biomarker of aging and disease. *Crit Rev Clin Lab Sci*. 2018;55:443–65.
- Blackburn EH. Telomere states and cell fates. *Nature*. 2000;408:53–6.
- Alzain MA, Asweto CO, Zhang J, Fang H, Zhao Z, Guo X, Song M, Zhou Y, Chang N, Wang Y. Telomere length and accelerated biological aging in the china suboptimal health cohort: a case-control study. *Omic*. 2017;21:333.
- Yeh JK, Wang CY. Telomeres and telomerase in cardiovascular diseases. *Genes*. 2016;7:58.
- Butt HZ, Atturu G, London NJ, Sayers RD, Bown MJ. Telomere length dynamics in vascular disease: a review. *Eur J Vasc Endovasc Surg*. 2010;40:17–26.
- Tian Y, Wang S, Jiao F, Kong Q, Liu C, Wu Y. Telomere length: a potential biomarker for the risk and prognosis of stroke. *Front Neurol*. 2019;10:624.
- Gao D, Zhang R, Ji G, Li C, Guo D, Jin T, Chen M. Relative telomere length and stroke risk in a chinese han population. *J Mol Neurosci*. 2018;2018:475–81.
- Luo D, Hou Q, Yu J, Yu D. Telomere length associated with the risks of high-risk and ischemic stroke in southern Chinese Han population. *Oncotarget*. 2017;8:105915–22.
- Ding H, Chen C, Shaffer JR, Liu L, Xu Y, Wang X, Hui R, Wang DW. Telomere length and risk of stroke in Chinese. *Stroke*. 2012;43:658–63.
- Schurks M, Prescott J, Dushkes R, De Vivo I, Rexrode KM. Telomere length and ischaemic stroke in women: a nested case-control study. *Eur J Neurol*. 2013;20:1068–74.
- Zee RY, Castonguay AJ, Barton NS, Ridker PM. Relative leukocyte telomere length and risk of incident ischemic stroke in men: a prospective, nested case-control approach. *Rejuvenation Res*. 2010;13:411–4.
- Ellehoj H, Bendix L, Osler M. Leucocyte telomere length and risk of cardiovascular disease in a cohort of 1397 Danish men and women. *Cardiology*. 2016;133:173–7.
- Jin X, Pan B, Dang X, Wu H, Xu D. Relationship between short telomere length and stroke: a meta-analysis. *Medicine (Baltimore)*. 2018;97:e12489.
- Haycock PC, Heydon EE, Kaptoge S, Butterworth AS, Thompson A, Williet P. Leucocyte telomere length and risk of cardiovascular disease: systematic review and meta-analysis. *BMJ-BRIT Med J*. 2014;349:g4227.

20. Sekula P, Del Greco MF, Pattaro C, Kottgen A. Mendelian randomization as an approach to assess causality using observational data. *J Am Soc Nephrol*. 2016;27:3253–65.
21. Hamad R, Walter S, Rehkopf DH. Telomere length and health outcomes: a two-sample genetic instrumental variables analysis. *Exp Gerontol*. 2016;82:88–94.
22. Cao W, Li X, Zhang X, Zhang J, Sun Q, Xu X, Sun M, Tian Q, Li Q, Wang H, et al. No causal effect of telomere length on ischemic stroke and its subtypes: a mendelian randomization study. *Cells*. 2019;8:159.
23. Haycock PC, Burgess S, Nounu A, Zheng J, Okoli GN, Bowden J, Wade KH, Timpson NJ, Evans DM, Willeit P, et al. Association between telomere length and risk of cancer and non-neoplastic diseases: a mendelian randomization study. *JAMA Oncol*. 2017;3:636–51.
24. Hernan MA, Robins JM. Instruments for causal inference: an epidemiologist's dream? *Epidemiology*. 2006;17:360–72.
25. Staley JR, Burgess S. Semiparametric methods for estimation of a nonlinear exposure-outcome relationship using instrumental variables with application to Mendelian randomization. *Genet Epidemiol*. 2017;41:341–52.
26. Stroke K. Recommendations on stroke prevention, diagnosis, and therapy. Report of the WHO task force on stroke and other cerebrovascular disorders. *Stroke*. 1989;20:1407–31.
27. Demanelis K, Jasmine F, Chen LS, Chernoff M, Tong L, Delgado D, Zhang C, Shinkle J, Sabarinathan M, Lin H, et al. Determinants of telomere length across human tissues. *Science*. 2020;369:6876.
28. O'Callaghan NJ, Fenech M. A quantitative PCR method for measuring absolute telomere length. *Biol Proced Online*. 2011;13:3.
29. O'Callaghan N, Dhillon V, Thomas P, Fenech M. A quantitative real-time PCR method for absolute telomere length. *Biotechniques*. 2008;44:807–9.
30. Fitzpatrick AL, Kronmal RA, Kimura M, Gardner JP, Psaty BM, Jenny NS, Tracy RP, Hardikar S, Aviv A. Leukocyte telomere length and mortality in the Cardiovascular Health Study. *J Gerontol A Biol Sci Med Sci*. 2011;66:421–9.
31. Weischer M, Bojesen SE, Nordestgaard BG. Telomere shortening unrelated to smoking, body weight, physical activity, and alcohol intake: 4576 general population individuals with repeat measurements 10 years apart. *PLoS Genet*. 2014;10:e1004191.
32. Jiang X, Dong M, Cheng J, Huang S, He Y, Ma K, Tang B, Guo Y. Decreased leukocyte telomere length (LTL) is associated with stroke but unlikely to be causative. *PLoS ONE*. 2013;8:e68254.
33. Li J, Feng C, Li L, Yang S, Chen Y, Hui R, Zhang M, Zhang W. The association of telomere attrition with first-onset stroke in Southern Chinese: a case-control study and meta-analysis. *Sci Rep*. 2018;8:2290.
34. D'Mello MJ, Ross SA, Briel M, Anand SS, Gerstein H, Pare G. Association between shortened leukocyte telomere length and cardiometabolic outcomes: systematic review and meta-analysis. *Circ-Cardiovasc Genet*. 2015;8:82–90.
35. Davies NM, Holmes MV, Davey SG. Reading Mendelian randomisation studies: a guide, glossary, and checklist for clinicians. *BMJ-BRIT Med J*. 2018;362:k601.
36. Davey Smith G, Hemani G. Mendelian randomization: genetic anchors for causal inference in epidemiological studies. *Hum Mol Genet*. 2014;23:R89–98.
37. Brion MJ, Shakhbazov K, Visscher PM. Calculating statistical power in Mendelian randomization studies. *Int J Epidemiol*. 2013;42:1497–501.

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